

Working Paper No. 1008

A GARCH Approach to Modeling Chilean Long-Term Swap Yields

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May 2022

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ISSN 1547-366X

ABSTRACT

This paper econometrically models the dynamics of the Chilean interbank swap yields based on macroeconomic factors. It examines whether the month-over-month change in the short-term interest rate has a decisive influence on the long-term swap yield after controlling for other factors, such as the change in inflation, change in the growth of industrial production, change in the log of the equity price index, and change in the log of the exchange rate. It applies the generalized autoregressive conditional heteroskedasticity (GARCH) approach to model the dynamics of the long-term swap yield. The change in the short-term interest rate has an economically meaningful and statistically significant effect on the change of the interbank swap yield. This means that the Banco Central de Chile's (BCCH) monetary policy exerts an important influence on interbank swap yields in Chile.

KEYWORDS: Interest Rate Swaps; Swap Yield; Short-Term Interest Rate; Banco Central de Chile (BCCH); Chile

JEL CLASSIFICATIONS: E43; E50; E58; E60; G10; G12

I. INTRODUCTION

Interbank interest rate swaps are an important component of the global over the counter (OTC) derivatives market. The notional value of outstanding interest rate swaps amounts to more than \$370 trillion¹ as of the second half of 2021, while the gross market value of interest rate swaps was \$8 trillion during the same period, according to the Bank for International Settlements (BIS) (2022). Interest rate swaps constitute more than 60 percent of over the OTC outstanding derivatives by notional value and almost 64 percent of outstanding derivatives by the gross market value. Yet, careful and detailed empirical analysis of the determinants of interbank swap yields has been limited. Even though there is considerable literature on swaps, there is a dearth of empirical modeling of interbank swap yields not just for emerging markets, such as Chile, but also for the interbank interest rate swap yields in advanced countries.

Interest rate swaps are increasingly important for financial markets and financial institutions even in emerging markets. While the bulk of interest rate swaps are denominated in the major currencies, such as the US dollar, the euro, the British pound, and the Japanese yen, the amount of outstanding interest rate swaps in other currencies, including emerging market currencies, is still substantial. As of the second half of 2021, BIS (2022) reports that for nonmajor currencies the notional amount of outstanding interest rate swaps is \$60 trillion and the gross market value is over \$720 billion.

As financial markets develop in emerging markets and these emerging markets undergo financialization, the analysis of the dynamics of interbank swap yields in emerging markets shall warrant vigilant attention and econometric analysis. This paper initiates this inquiry. It fills a consequential gap in the empirical literature regarding the macroeconomic determinants of interbank swap yields. The dynamics of Chilean long-term swap yields are analyzed in this paper through examining whether the month-over-month change in the short-term interest rate has an influence on the month-over-month change in the long-term swap yields after controlling for other factors, such as the change in inflation, the change in the growth of industrial production, the change in the log of the equity price index, and the change in the log of the exchange rate.

¹ All figures are in US dollars unless specified otherwise.

This paper is arranged as follows. Section II explains what an interest rate swap is. It also briefly reviews the literature on interest rate swaps. Section III presents a simple model that connects the interbank swap yield to the short-term interest rate and other macroeconomic factors. Section IV provides the macroeconomic backdrop to the evolution of the interbank swap yield in Chile. Section V gives the data sources of the variables used in the econometric modeling of the swap yield, explains these variables, and undertakes unit root and stationarity tests. Section VI reports the findings from the econometric modeling of the interbank swap yield. Section VIII briefly discusses the implications of these findings. Section VIII concludes.

II. INTEREST RATE SWAPS AND A BRIEF REVIEW OF THE LITERATURE

A swap is a type of a financial contract. In an interbank interest rate swap contract, two parties exchange cash flows with different characteristics. Usually two banks (and/or other any financial institutions) exchange cash flows on two different types of interest payments. The principal amount is the same for both banks. This is known as the notional principal. One bank typically pays a fixed interest rate on the principal amount to the other bank, while in return it a receives a variable interest rate from the other bank. The bank that *receives* the variable interest rate buys the interbank swap, whereas the bank that *pays* the variable interest rate sells the interbank swap (figure 1).

Figure 1. The Bank That Buys the Interbank Interest Rate Swap Makes a Payment Based on a Fixed Interest Rate and Receives a Payment Based on a Variable Interest Rate



If a bank expects the interest rate to rise it would buy the swap because it would lock in the amount that it would have to pay in exchange for the variable interest rate payments that it would receive from the other bank. Likewise, if a bank expects the interest rate to decline, it will sell the swap because it would lock in the amount that it would receive in exchange for the variable interest rate payments it would pay to the other bank.

The fixed interest rate payment of the swap is paid semiannually for the maturity tenor of the swap. This is known as the *swap yield* or the *swap rate* for the tenor under consideration. The variable interest rate payment is linked to some benchmark interest rate. As the benchmark interest rate changes, the variable interest rate also changes. The variable payments are calculated based on the variable interest rate. In each quarter the benchmark interest rate is registered to determine the variable interest rate and the variable interest payment. The variable interest payment is made at the end of the quarter. The present value of the fixed and variable legs of the swap are the same at its inception. Swaps are conducted among the contracting parties over the counter rather than on a financial exchange. The tenor of the swaps can vary, ranging from overnight to over 30 years.

There is substantial literature on interest rate swaps. Bicksler and Chen (1986) give an economic analysis of interest rate swaps and their use in finance and business. They describe alternative uses of and the appropriate valuation procedure for interest rate swaps. Corb (2012) provides a broad overview, explains the concepts behind interest rate swaps, and explores key themes concerning swaps, such as their risk characteristics, traditional use, and pricing, as well as swaptions and recent innovations in swaps. Remolona and Wooldridge (2003) survey eurodenominated interest rate swaps. They examine the size of the euro swap market, the role of swaps as benchmark instruments, and the pricing of swaps. Chernenko and Faulkender (2011) canvass firms on the use of interest rate swaps. They report that hedging of interest rate risk is concentrated among high-investment firms. They also find that firms appear to use interest rate swaps to manage earnings and sometimes to engage in speculation. Duffie and Huang (1996) develop a model that relates the credit quality of a corporation to the swap yield. Kim and Koppenhaver (1993) find that the likelihood and extent of swap market participation by low-capitalized banks is less than for other banks. Visvanathan (1998) finds that firms that expect

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high financial distress costs use swaps to transform short-term debt into long-term fixed-rate debt. Debt maturity structure is significant in the decision to use a swap. Empirical research on swaps, such as Lekkos and Milas (2001), has been confined to relating the swap yield to business cycle conditions rather than fundamental macroeconomic and financial variables. Duffie and Singleton (1997) develop a multifactor econometric model of the term structure of interest rate swap yields. They report that both credit and liquidity factors are crucial drivers of the swap yield, but they too do not analyze the macro dynamics of the swap yield. It is apparent that the scholarly literature on interest rate and the long-term swap yield have not been explored in the finance literature.

The relationship between the short-term interest rate and the long-term government bond yield, which has been thoroughly investigated, provides a useful basis for examining the dynamics of the long-term swap yield from a macroeconomic vantage point and filling a consequential gap in the literature. Keynes (1930, [1936] 2007) maintains that the central bank's actions have decisive effects on the long-term government bond yield, primarily through the influence of the policy rate on the short-term interest rate. Keynes's conjecture about this relationship drew upon Riefler's (1930) inference, which came from detailed statistical analysis of interest rate dynamics in the 1920s in the United States and Keynes's own observations about interest rate dynamics in the United Kingdom during the same period. Kregel (2011) explores and reprises Keynes's views on the influence of the central bank's policy rate on long-term government bonds yields, investors' behavior in financial behavior, and fundamental uncertainty.

Recent empirical research on long-term government bonds yields, such as Akram and Li (2020a, 2020b, 2020c), has bolstered support for the conjecture that the short-term interest rate is a key driver of the long-term government bond yield. Moreover, these researchers and others show that the change in the short-term interest rate is a key driver of the change in the long-term government bond yield. Keynes's conjecture that relates monetary policy actions to the dynamics of the long-term government bond yield provides a fecund theoretical and empirical basis for modeling the swap yield as a function of the short-term interest rate and examining whether there

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is an empirical relationship between the short-term interest rate and the swap yield after controlling for relevant macroeconomic and financial variables.

III. A MODEL OF THE INTERBANK SWAP YIELD

A model of the interbank swap yield is presented here. Akram's (2021, 2022) models operationalize Keynes's insight that the short-term interest rate is the primary driver of the long-term government bond yield. The model presented here modifies Akram's (2021, 2022) models to make them suitable for analyzing the dynamics of the long-term swap yield.

The long-term interbank swap yield is S_{LT} . The short-term interest rate is i_{ST} . The central bank's policy rate is i_{CB} . The inflation is π , while the central bank's inflation target is $\overline{\pi}$. χ represents financial market volatility, while $\tau(t)$ is an exogenous shock. W(t) is the Weiner process. The parameters of the models are: $\alpha_1, \alpha_2, \beta, \gamma, \delta$.

$$dS_{LT}(t) = \left(\alpha_1 i_{ST}(t) + \alpha_2 \pi(t)\right) dt + \chi(t) \sqrt{i_{ST}(t)} dW(t)$$
^[1]

$$di_{ST}(t) = \beta(i_{CB}(t) - i_{ST}(t))dt + \chi(t)\sqrt{i_{ST}(t)}dW(t)$$
[2]

$$d\pi(t) = \gamma(\bar{\pi} - \pi(t))dt + \chi(t)\sqrt{\pi(t)}dW(t)$$
[3]

$$d\chi(t) = \delta(\bar{\chi} - \chi(t))dt + \tau(t)\sqrt{\chi(t)}dW(t)$$
[4]

Equation [1] relates the dynamics of the long-term swap yield to the change in the short-term interest rate, the change in inflation, and the change in the Weiner process adjusted by the volatility of financial markets and the short-term interest rate. Equation [2] expresses the dynamics of the short-term interest rate as a function of (1) the difference between the central bank's policy rate and the short-term interest rate, and (2) the Weiner process adjusted by the volatility of the financial market and the short-term interest rate. Equation [3] relates the

dynamics of inflation to (1) the difference between the central bank's inflation target and inflation, and (2) the Weiner process adjusted by the volatility of financial market and inflation. Equation [4] relates the dynamics of the financial market's volatility to a mean reverting process and the Weiner process adjusted by an exogenous shock and the volatility of the financial market.

The above model ties the dynamics of the interbank swap yield to fundamental macroeconomic and financial variables, such as the change in the short-term interest rate, change in inflation, and financial market volatility. It can be seamlessly extended to incorporate any other pertinent macroeconomic factor, such as the change in the growth of industrial production, change in the logarithm of the equity price index, and change in the logarithm of the exchange rate, if these factors are deemed as important drivers of the interbank swap yield.

Later in this paper the standard GARCH(1,1) approach is applied to econometrically model the dynamics of the swap yield and relate it to the change in the short-term interest rate and other macroeconomic and financial variables.

IV. MACROECONOMIC BACKDROP TO THE EVOLUTION OF THE INTERBANK SWAP YIELD IN CHILE

Even a perfunctory analysis of the stylized facts of macroeconomic and financial data would reveal that monetary policy and overall interest rate dynamics have a profound influence on the change in the swap yield. Chile's interbank swap yield follow similar patterns.

Figure 2 shows the evolution of the interbank swap yield and the short-term interest rate in Chile between 2005 and 2021. Between 2005 and 2007, the interbank swap yield on swaps of different maturity tenors steadily increased as the Banco Central de Chile (BCCH) raised its policy rate. The interbank swap yield declined sharply during the global financial crisis (GFC) as the shortterm interest rate declined in lockstep with the BCCH's policy rate. The interbank swap yield gradually rose from early 2009 to mid-2011. From mid-2011 to 2019 the interbank swap interest

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yield gradually declined. As BCCH cut its policy rate in response to the global lockdown during the COVID-19 pandemic, the interbank swap yield fell markedly. Since mid-2021 the interbank swap yield rose noticeably as the BCCH raised its policy rate.



Figure 2. The Evolution of Interest Rate Swap Yields and the Short-term Interest Rate in Chile, 2005–21

Figure 3 displays the coevolution of the interbank swap yield and consumer price index (CPI) inflation in Chile. The swap yield and inflation generally appear to move together, though the relationship between the swap yield and inflation is often rather weak.



Figure 3. The Coevolution of 10-year Interbank Swap Yield and CPI Inflation, 2005–21

Figure 4 shows the evolution of inflation and core inflation in Chile from 2005 to 2021. Overall inflation and core inflation tend to move together. Inflation rose sharply in mid-2007 and continued to rise until later in 2008. Inflation fell between 2009–10. Except for a brief spell of high inflation between late 2014 and early 2015, inflation stayed in the range of 2–4 percent year over year between mid-2010 and mid-2021. Inflation began rising in mid-2021. By late 2021, inflation exceed 6 percent, while core inflation was just shy of 6 percent.



Figure 5 displays the growth of industrial production. The growth of industrial production is a useful indicator of business cycle conditions of economic activity in Chile. The time series on the growth of industrial production is volatile, but it shows that industrial production usually trends to grow. However, industrial production declined during the GFC. Industrial production rose in the subsequent quarters, exhibiting recovery from the GFC. However, between early 2013 to late 2019 the country's industrial production exhibited considerable volatility from month to month. During the global lockdown, industrial production declined sharply amid restrictions and social distancing but its growth resumed as restrictions were later scaled back and the pandemic subsided.





Figure 6 exhibits the evolution of the Chilean peso against the US dollar. The Chilean peso appreciated more than 20 percent between January 2005 to mid-2008. It subsequently depreciated sharply about 40 percent between mid-2008 and early 2009. This depreciation was reversed between 2009 and early 2011. The peso remained steady for the next two years. The peso depreciated between early 2013 to early 2016, followed by moderate appreciation until early 2018. It depreciated from April 2018 to April 2020, followed by appreciation until May 2021. However, the peso soon reversed course; it depreciated from mid-2021 to the end of the year.



Figure 6. The Evolution of the Chilean Peso, USDCLP, 2005–21

Figure 7 displays the evolution of the Chilean equity price index, as measured by the Indice General de Precios de Acciones (IGPA), between 2005–21. It was generally rising throughout the period, though there were some periods during which IGPA either declined or stayed flat. For instance, between mid-2007 until late 2008 it declined. It was range bound between mid-2011 and mid-2016. After climbing from early 2016 to early 2018, the index declined until mid-2020. It was range bound from 2020 until the end of 2021.



Figure 7. The Evolution of the Chilean Equity Price Index, IGPA, 2005–21

V. DATA DESCRIPTION, UNIT ROOT TESTS, AND STATIONARITY TESTS

Table 1 below provides a summary of the variables used in the paper. The first column lists the variable names. The second column gives the data description and the date range for the data. The third column enumerates the data frequency and indicates whether high frequency data have been converted to low frequency data. The final column provides the source of the data.

Two different variables are used for the short-term interest rate. These short-term interest rates are based on the Pagaré Descontable del Banco Central de Chile (PDBC) instruments, which are discountable promissory notes issued by the BCCH, the nation's central bank. There are three

different interbank swap yields in two-year, five-year, and ten-year maturity tenors. Inflation is measured using both total inflation and core inflation. Total inflation is based on the total CPI, while core inflation is the total CPI excluding food and energy prices. Economic activity is measured by the growth of industrial production year over year. The exchange rate is based on the value of Chilean peso per US dollar. IGPA general is the equity price index.

Monthly data for the above-mentioned variables are used. For all variables (except core inflation) the data's time range is from January 2005 to December 2021, consisting of more than 200 observations.

Variables	Data description, date range	Frequency	Sources
Short-term interest rate	25		
PDBC30D	Interest rate on BCCH instrument, PDBC 30 days, %, Jan 2005–Dec 2021	Daily; converted to monthly	Banco Central de Chile
PDBC90D	Interest rate on BCCH instrument, PDBC 90 days, %, Jan 2005–Dec 2021	Daily; converted to monthly	Banco Central de Chile
Long-term swap rates			
SWAP2Y	Interbank swap yield, 2 year, %, January 2005–December 2021	Daily; converted to monthly	Banco Central de Chile
SWAP5Y	Interbank swap yield, 5 year, %, January 2005–December 2021	Daily; converted to monthly	Banco Central de Chile
SWAP10Y	Interbank swap yield, 10 year, %, January 2005–December 2021	Daily; converted to monthly	Banco Central de Chile
Inflation			
СРІҮОҮ	Consumer price index, all items, seasonally adjusted, 2018 = 100, % change, y/y, January 2005–December 2021	Monthly	Instituto Nacional de Estadística de Chile
COREYOY	Consumer price index, all items excluding food and energy, seasonally adjusted, 2018 = 100, % change, y/y, January 2010– December 2021	Monthly	Instituto Nacional de Estadística de Chile
Economic activity			
ΙΡΥΟΥ	Industrial production index, seasonally adjusted, 2014 = 100, % change, y/y, January 2005–December 2021	Monthly	Sociedad de Fomento Fabril
Financial variables			
CLP	Exchange rate, Chilean peso per U.S. dollar, USDCLP, January 2005–December 2021	Daily; converted to monthly	Banco Central de Chile
IGPA	Equity price index, IGPA General, 12/31/1980 = 100, January 2005–December 2021	Daily; converted to monthly	Financial Times

Table 1. Summary of the Data

Note that in the text below, LNIGPA indicates the (natural) logarithm of IGPA. Likewise, LNCLP is the (natural) logarithm of CLP, the exchange rate.

Table 2A and table 2B provide the summary statistics of these variables. Table 2A displays the summary statistics of the variables, while table 2B provides the summary statistics of the first differences of the same variables. Table 2A shows that most variables, except for CPIYOY, IPYOY, LNIGPA, and LNCLP, are not normally distributed. However, table 2B reveals that the first differences of all the variables are all normally distributed.

	Obs	Mean	Std. Dev.	Max	Min	Skewness	Kurtosis	J-B	Probability
SWAP2Y	204	4.03	1.62	8.30	0.54	-0.01	2.74	0.56	0.76
SWAP5Y	204	4.65	1.43	7.94	1.17	-0.32	2.89	3.52	0.17
SWAP10Y	204	5.15	1.26	7.82	2.16	-0.31	2.67	4.27	0.12
PDBC30D	204	3.43	1.80	8.74	0.24	0.27	3.00	2.55	0.28
PDBC90D	204	3.51	1.80	8.36	0.24	0.18	2.80	1.47	0.48
СРІУОУ	204	3.42	2.03	9.90	-2.22	0.79	5.00	55.28	0.00
COREYOY	144	2.48	1.34	6.42	-1.70	0.39	3.23	3.97	0.14
ΙΡΥΟΥ	204	1.28	5.95	30.53	-17.84	0.32	5.81	70.61	0.00
LNIGPA	204	9.80	0.30	10.27	9.08	-0.79	2.73	21.57	0.00
LNCLP	204	6.38	0.16	6.75	6.09	0.37	2.06	12.13	0.00

Table 2A. Summary Statistics of the Variables

Table 2B. Summary statistics of the first differences of the variables

	Obs	Mean	Std. Dev.	Max	Min	Skewness	Kurtosis	J-B	Probability
∆SWAP2Y	203	0.01	1.09	-1.34	0.30	-0.77	8.12	242.12	0.00
∆SWAP5Y	203	0.00	0.96	-1.02	0.26	-0.21	5.88	71.76	0.00
∆SWAP10Y	203	-0.01	0.88	-0.96	0.23	-0.02	5.50	52.84	0.00
∆PDBC30D	203	0.01	1.26	-2.66	0.37	-2.18	19.47	2454.84	0.00
∆PDBC90D	203	0.01	1.50	-2.31	0.37	-1.76	16.83	1722.40	0.00
ΔCPIYOY	203	0.02	1.47	-1.92	0.51	-0.63	4.83	41.68	0.00
ACOREYOY	143	0.06	0.32	1.82	0.65	1.30	8.58	225.62	0.00
ΔΙΡΥΟΥ	203	-0.01	26.64	-22.41	5.85	0.24	5.52	55.68	0.00
ΔLNIGPA	203	0.00	0.11	-0.23	0.04	-1.23	8.44	301.45	0.00
ΔLNCLP	203	0.00	0.15	-0.07	0.03	1.04	7.84	235.17	0.00

The unit root tests are conducted using the automated Dickey-Fuller (ADF) tests (Dickey and Fuller 1979, 1981), while the stationarity tests are conducted using the Kwiatkowski, Phillips, Schmidt, and Shin (KPSS) tests (Kwiatkowski et al. 1992).

Table 3A displays the results of the unit root tests and the stationary tests for these variables. The results are mixed regarding swap yields, while PDBC rates are stationary with the intercept and trend inclusion in the test equation for an ADF unit root test but nonstationary otherwise. Similar mixed results are obtained in the KPSS stationary tests. Among the control variables, only the growth of industrial production yielded a stationary result in both types of tests. Other control variables are either nonstationary or show mixed results under different assumptions.

	ADF Unit	Root Tests (Ho: Nonst	ationary)	KPSS Tests (H ₀ : Stationarity) tests			
	None	Intercept	Trend	Intercept	Trend		
SWAP2Y	-0.79	-2.57	- 3.49**	1.11***	0.05		
SWAP5Y	-0.62	- 2.16	- 3.61**	1.42***	0.05		
SWAP10Y	- 0.66	- 2.13	- 3.88**	1.52***	0.05		
PDBC30D	-1.00	-2.67*	- 2.96*	0.60**	0.08		
PDBC90D	- 1.17	- 3.21**	- 3.70**	0.65**	0.05		
CPIYOY	-0.53	-2.38	-2.27	0.10	0.05		
COREYOY	0.82	- 0.99	- 1.26	0.37	0.20**		
IPYOY	- 3.38***	-3.50***	- 3.50**	0.07	0.07		
LNIGPA	-1.47	- 2.39	-1.78	1.32***	0.24***		
LNCLP	-0.79	-0.67	- 2.51	1.27***	0.28***		

Table 3A. Unit Root and Stationarity Tests of the Variables

Note: Significance levels *** for 1 percent, ** for 5 percent, and * for 10 percent

Table 3B presents the unit root and the stationarity tests for the first differences of the variables. All the ADF unit root tests indicate the null hypothesis on nonstationarity can be rejected at the 1 percent level of significance for the first differences of all variables. The KPSS tests show that the null hypothesis of stationarity cannot be rejected for the first differences of these variables (except for the growth of industrial production at 10 percent significance under trend inclusion).

	ADF Uni	t Root Test (Ho: Nonst	ationary)	KPSS Test (H ₀ : Stationarity)			
	None	Intercept	Trend	Intercept	Trend		
∆SWAP2Y	-7.11***	- 7.09***	-7.09***	0.08	0.07		
∆SWAP5Y	-8.43***	-8.41***	-8.40***	0.09	0.08		
∆SWAP10Y	-8.96***	-8.94***	-8.93***	0.08	0.06		
∆PDBC30D	- 7.27***	- 7.24***	-7.24***	0.05	0.05		
∆PDBC90D	- 5.31***	- 5.30***	-5.28***	0.06	0.06		
ΔCPIYOY	-6.06***	-6.06***	-6.06***	0.06	0.05		
<i>ACOREYOY</i>	- 5.90***	-6.09***	-6.07***	0.20	0.19		
ΔΙΡΥΟΥ	- 15.92***	-15.88***	-10.15***	0.45	0.50*		
ΔLNIGPA	- 12.16***	- 12.25***	- 12.36***	0.26	0.05		
ALNCLP	-10.28***	-10.31***	-10.39***	0.19	0.03		

Table 3B. Unit Root and Stationarity Tests of the First Differences of the Variables

Note: Significance levels *** for 1 percent, ** for 5 percent, and * for 10 percent

The unit root tests and the stationary tests imply that it is appropriate to econometrically model the month-over-month change in the swap yield using month-over-month change in the shortterm interest rate, change in inflation, change in the growth of industrial production, change in the log of the equity price index, and change in the exchange rate.

VI. ECONOMETRIC MODELS AND EMPIRICAL RESULTS

The autoregressive conditional heteroskedasticity (ARCH) Lagrange multiplier (LM) tests on ordinary least square (OLS) regressions of swap yield models are conducted to ascertain whether an ARCH framework is a suitable approach for econometrically modeling the dynamics of the swap yield. These models and their generalized version (GACRH) are specifically designed to model and forecast conditional variances.

ARCH models were introduced by Engle (1982) and GARCH by Bollerslev (1986) and Taylor (1986).² In ARCH and GARCH models, the variance of the dependent variable is a function of the past values of the dependent variable and independent, or exogenous, variables. This allows the analyst to model volatility over time.

² For additional background information, including the econometric theory and some applications, see Bollerslev, Chou, and Kroner (1992) and Bollerslev, Engle, and Nelson (1994). These two papers provide comprehensive surveys of ACRH and GARCH models and their applications.

The ARCH LM tests are given in table 4. The tests show that the presence of ARCH in the OLS regression models of the month-over-month change in the swap yield of different maturity tenors. These results clearly indicate that an ARCH-type model will be useful for estimating the relationship between the month-over-month change in the swap yield and the month-over-month change in the short-term interest rate in Chile, after controlling for other factors, by modeling the volatility.

Models	ΔSWAP2Y	∆SWAP5Y	∆SWAP10Y	ΔSWAP2Y	ΔSWAP5Y	∆SWAP10Y		
Lags		∆PDBC30D		∆PDBC90D				
1	9.23	7.25	13.85	15.71	9.91	13.38		
	(0.00)	(0.01)	(0.00)	(0.00)	(0.00)	(0.00)		
4	5.14	5.77	6.31	7.73	5.83	6.60		
	(0.00)	(0.00)	(0.00)	(0.03)	(0.00)	(0.00)		
8	2.72	4.05	3.96	4.23	4.33	4.00		
	(0.01)	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)		
12	1.91	2.73	2.84	2.89	3.23	2.96		
	(0.04)	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)		

Table 4. ARCH LM Test

Note: OLS model includes the change in the short-term interest rate (Δ PDBC30D, Δ PDBC9oD) and the controls (namely Δ CPIYOY, Δ IPYOY, Δ LNIGPA, and Δ LNCLP). *p*-values are in parenthesis.

To address these issues and allow the conditional variance of the error term to depend upon its previous own lags, the following standard GARCH(1,1) model is used here to econometrically analyze the dynamic of the swap yield.

$$Y_t = \beta_0 + \beta_1 X_t + \varepsilon_t \tag{5}$$

$$\sigma_t^2 = \alpha_0 + \alpha_1 \varepsilon_{t-1}^2 + \gamma_1 \sigma_{t-1}^2$$
[6]

Here the current volatility of the error term is explained by the long-run average variance (α_0), the past values of the shocks, and the history of volatility.

The following GARCH(1,1) models are estimated as specified below:

$$\Delta SWAP2Y = \varphi^{1}(C, \Delta PDBC90D, AR(1))$$

$$[7]$$

$$\Delta SWAP2Y = \varphi^{2}(C, \Delta PDBC90D, \Delta CPIYOY, \Delta IPYOY, AR(1))$$

$$[8]$$

$$\Delta SWAP2Y = \varphi^{3}(C, \Delta PDBC90D, \Delta CPIYOY, \Delta IPYOY, \Delta LNIGPA, \Delta LNCLP, AR(1))$$

$$[9]$$

$$\Delta SWAP5Y = \varphi^{4}(C, \Delta PDBC90D, AR(1))$$

$$[10]$$

$$\Delta SWAP5Y = \varphi^{5}(C, \Delta PDBC90D, \Delta CPIYOY, \Delta IPYOY, AR(1))$$

$$[11]$$

$$\Delta SWAP5Y = \varphi^{6}(C, \Delta PDBC90D, \Delta CPIYOY, \Delta IPYOY, ALNIGPA, \Delta LNCLP, AR(1))$$

$$[13]$$

$$\Delta SWAP10Y = \varphi^{9}(C, \Delta PDBC90D, \Delta CPIYOY, \Delta IPYOY, AR(1))$$

$$[14]$$

$$\Delta SWAP10Y = \varphi^{9}(C, \Delta PDBC90D, \Delta CPIYOY, \Delta IPYOY, AR(1))$$

$$[15]$$

$$\Delta SWAP2Y = \psi^{1}(C, \Delta PDBC30D, AR(1))$$

$$[16]$$

$$\Delta SWAP2Y = \psi^{2}(C, \Delta PDBC30D, \Delta CPIYOY, \Delta IPYOY, AR(1))$$

$$[17]$$

$$\Delta SWAP5Y = \psi^4(C, \Delta PDBC30D, AR(1))$$
[19]

$$\Delta SWAP5Y = \psi^{5}(C, \Delta PDBC30D, \Delta CPIYOY, \Delta IPYOY, AR(1))$$
[20]

 $\Delta SWAP5Y = \psi^{6}(C, \Delta PDBC30D, \Delta CPIYOY, \Delta IPYOY, \Delta LNIGPA, \Delta LNCLP, AR(1))$ [21]

$$\Delta SWAP10Y = \psi^{7}(C, \Delta PDBC30D, AR(1))$$
[22]

$$\Delta SWAP10Y = \psi^{8}(C, \Delta PDBC30D, \Delta CPIYOY, \Delta IPYOY, AR(1))$$
[23]

Δ SWAP10Y = ψ^9 (C, Δ PDBC30D, Δ CPIYOY, Δ IPYOY, Δ LNIGPA, Δ LNCLP, AR(1)) [24]

The main results for the GARCH(1,1) models are presented in table 5A and table 5B using Δ PDBC90D and Δ PDBC30D respectively as the month-over-month change in the short-term interest rate. The swap rates for two, five and ten years are modeled by adding various control variables (namely Δ CPIYOY, Δ IPYOY, Δ LNIGPA, and Δ LNCLP). An autoregressive term (AR(1)) is added to control for the autocorrelation in the models. Model diagnostic information and the results of postestimation diagnostic tests are also displayed in these tables. The correlograms (autocorrelations and partial autocorrelations) of the standardized residuals from the estimated GARCH(1,1) models with Δ PDBC90D and Δ PDBC30D, respectively, are provided in appendix A and appendix B.

In the mean equation of table 5A, the effect of Δ PDBC90D on Δ SWAP2Y and Δ SWAP5Y is positive and statistically significant. However, its effect on Δ SWAP10Y is positive but not statistically significant. This means that the change in the short-term interest rate has clear, definitive, and statistically significant effects on the frontend and the belly of the interbank swap yield curve but not on the backend of the swap yield curve. The effects of Δ CPIYOY and Δ IPYOY on the swap yield of different maturity tenors are positive but mostly not statistically significant. Among the control variables, the change in the log of the Chilean equity price index has a positive and statistically significant effect on the change in the swap yield of all maturity tenors. This means as that a rise (fall) in the equity price index is associated with an increase (decline) in the swap yields of different maturities. The results also show the change in the log of the exchange rate of the Chilean peso against the US dollar has a positive and statistically significant effect on Δ SWAP5Y and Δ SWAP10Y. However, its effect on Δ SWAP2Y is negative but not statistically significant. The AR(1) term has positive and statistically significant effect on the swap yield of all three maturity tenors.

The parameters in the variance equation are statistically significant. The significant ARCH coefficient implies that a volatility shock today feeds into the next month's volatility. The significant GARCH coefficient indicates a large shock (either positive or negative) will lead to a large variance in the forecast for a long period of time. The sum of the ARCH and the GARCH coefficients measures the rate at which the volatility effect fades over time. Since the sum is high, the shocks to the conditional variance are persistent and clustered over time.

Table 5B shows that \triangle PDBC30D has pretty much the same effect as that of \triangle PDBC90D on the swap yield of different maturity tenors. The effect of \triangle PDBC30D on not just \triangle SWAP2Y and \triangle SWAP5Y but also on \triangle SWAP10Y is positive and statistically significant. However, the effect's magnitude on the change in the swap yields of longer maturity tenors is smaller. The effects of \triangle CPIYOY and \triangle IPYOY on the swap yield of different tenors are positive but mostly not statistically significant. The effect of \triangle LNIGPA on the swap yield is always positive and statistically significant while the effect of \triangle LNCLP on the swap yield is always positive and statistically significant. The effect of the AR(1) term is always positive and statistically significant.

In the models with Δ PDBC30D, the ARCH and the GACRH coefficients in the variance equation are both statistically significant. The sum of the two coefficients is closer to one than in the models presented earlier in table 5A. This indicates strong evidence of the persistence and clustering of the variance in the error terms. Here, too, the positive and statistically significant ARCH coefficient implies that a volatility shock today feeds into the next month's volatility. The positive and statistically significant GARCH coefficient indicates a large shock (either positive or negative) will lead to a large variance in the forecast for a long period of time.

It is useful to have some perspective on the financial market volatility that can affect the interbank swap yields and financial conditions in Chile. Chile is a high-income emerging market but it is subject to financial shocks, international trade slowdowns, and global economic pressures. Exports of goods and services account for nearly one-third of the country's nominal GDP. Commodities make up nearly 60 percent of Chile's total exports. Copper is the country's main export, providing around 20 percent of government revenue. Due to the high share of exports in nominal GDP and its dependence on the revenue from the export of copper, Chile's financial markets and its economy can exhibit volatility emanating from the uncertainty regarding fluctuations in commodity prices (especially the international price of copper) and global industrial production, as well as turbulence in overseas financial markets and economic shocks.

Postestimation tests for these models show support for the GARCH approach to the econometric modeling of the swap yield of various maturity tenors as evinced by the ARCH LM tests. The models do not have any autocorrelation problems and the standardized residuals are normally distributed. The correlograms in appendix A and appendix B show that there is no remaining autocorrelation in the mean equation and that the chosen models are correctly specified.

The models are re-estimated by replacing inflation with core inflation as a control variable. The results, which are displayed in appendix C, are comparable to the original models.

	∆SWAP2Y	∆SWAP2Y	∆SWAP2Y	∆SWAP5Y	∆SWAP5Y	∆SWAP5Y	∆SWAP10Y	∆SWAP10Y	∆SWAP10Y
				Mean Ec	uation				
Intercept	-0.01	-0.01	-0.01	-0.005	-0.01	-0.01	-0.003	-0.005	-0.01
-	(0.84)	(0.83)	(0.71)	(0.83)	(0.72)	(0.54)	(0.84)	(0.82)	(0.64)
APDBC90D	0.29	0.29	0.31	0.13	0.12	0.13	0.05	0.04	0.04
	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	(0.24)	(0.31)	(0.35)
ΔCPIYOY		0.02	0.02		0.03	0.02		0.02	0.01
		(0.61)	(0.57)		(0.31)	(0.53)		(0.48)	(0.81)
ΔΙΡΥΟΥ		0.004	0.003		0.003	0.002		0.002	0.002
		(0.07)	(0.14)		(0.25)	(0.35)		(0.20)	(0.18)
ΔLNIGPA			0.81			0.80			0.43
			(0.01)			(0.03)			(0.27)
ΔLNCLP			-0.18			0.98			1.55
			(0.73)			(0.05)			(0.00)
AR(1)	0.41	0.40	0.38	0.33	0.33	0.34	0.37	0.35	0.37
	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)
			r	Variance l	Equation	r			
Intercept	0.01	0.01	0.01	0.01	0.01	0.01	0.01	0.005	0.004
	(0.15)	(0.16)	(0.15)	(0.20)	(0.21)	(0.23)	(0.13)	(0.13)	(0.17)
ARCH	0.13	0.15	0.19	0.13	0.13	0.14	0.13	0.13	0.14
	(0.03)	(0.03)	(0.04)	(0.08)	(0.07)	(0.08)	(0.01)	(0.01)	(0.02)
GARCH	0.73	0.70	0.67	0.72	0.75	0.75	0.71	0.74	0.77
	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)
	1	1	1	Model Inf	ormation	1	1	1	1
Obs	202	202	202	202	202	202	202	202	202
Adj R ²	0.43	0.43	0.44	0.26	0.25	0.26	0.19	0.18	0.19
AIC	-0.14	- 0.13	-0.14	- 0.20	-0.19	-0.19	-0.35	-0.34	-0.36
	1	1	1	Diagnost	ic Tests	1	1	1	1
ARCH LM	0.76	0.88	1.06	0.93	0.85	0.67	0.67	0.54	0.42
(12 lags)	(0.69)	(0.57)	(0.40)	(0.51)	(0.60)	(0.77)	(0.78)	(0.88)	(0.95)
DW Stat	1.88	1.90	1.90	1.80	1.82	1.82	1.84	1.85	1.82
JQB	46.69	33.78	18.39	13.96	13.61	11.09	7.25	6.73	6.28
	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	(0.03)	(0.04)	(0.05)

Table 5A. GARCH (1,1) Model (with **APDBC90D**)

Note: all vars are in diff, *p*-values are in parenthesis

	∆SWAP2Y	∆SWAP2Y	∆SWAP2Y	∆SWAP5Y	∆SWAP5Y	∆SWAP5Y	∆SWAP10Y	∆SWAP10Y	∆SWAP10Y
				Mean Ec	quation				
Intercept	-0.003	-0.004	-0.01	-0.002	-0.005	-0.01	-0.004	-0.01	-0.01
	(0.89)	(0.86)	(0.75)	(0.92)	(0.82)	(0.56)	(0.85)	(0.76)	(0.64)
∆PDBC30D	0.38	0.38	0.38	0.23	0.23	0.25	0.14	0.14	0.04
	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	(0.35)
ΔCPIYOY		0.02	0.02		0.05	0.03		0.03	0.01
		(0.51)	(0.50)		(0.13)	(0.35)		(0.33)	(0.81)
ΔΙΡΥΟΥ		0.003	0.002		0.002	0.002		0.002	0.002
		(0.15)	(0.19)		(0.26)	(0.33)		(0.19)	(0.18)
ΔLNIGPA			0.79			0.82			0.43
			(0.01)			(0.03)			(0.27)
ΔLNCLP			0.40			1.42			1.55
			(0.50)			(0.02)			(0.00)
AR(1)	0.39	0.38	0.37	0.34	0.32	0.34	0.35	0.33	0.37
	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)
				Variance l	Equation				
Intercept	0.002	0.002	0.002	0.003	0.002	0.002	0.004	0.002	0.004
	(0.13)	(0.17)	(0.17)	(0.30)	(0.22)	(0.28)	(0.17)	(0.18)	(0.17)
ARCH	0.08	0.09	0.10	0.08	0.08	0.09	0.11	0.11	0.14
	(0.03)	(0.04)	(0.04)	(0.08)	(0.03)	(0.06)	(0.03)	(0.02)	(0.02)
GARCH	0.88	0.87	0.84	0.85	0.88	0.87	0.79	0.84	0.77
	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)
				Model Info	ormation	-	-		
Obs	202	202	202	202	202	202	202	202	202
Adj R ²	0.47	0.47	0.48	0.31	0.30	0.31	0.22	0.21	0.19
AIC	-0.21	- 0.20	- 0.21	- 0.24	- 0.23	- 0.25	-0.37	-0.37	- 0.36
				Diagnost	ic Tests				
ARCH LM	1.46	1.56	1.09	0.46	0.49	0.34	0.51	0.41	0.67
(12 lags)	(0.14)	(0.11)	(0.37)	(0.93)	(0.92)	(0.98)	(0.90)	(0.96)	(0.79)
DW Stat	1.88	1.90	1.88	1.80	1.82	1.80	1.82	1.81	1.82
JQB	9.36	8.59	4.60	12.60	11.72	7.22	4.95	4.04	6.28
_	(0.01)	(0.01)	(0.10)	(0.00)	(0.00)	(0.03)	(0.08)	(0.13)	(0.05)

Table 5B. GARCH (1,1) Model (with △PDBC30D)

Note: all vars are in diff, *p*-values are in parenthesis

VII. THE IMPLICATIONS OF THE EMPIRICAL FINDINGS

The empirical findings reported in this paper have consequential implications for macroeconomic and financial theory, monetary policy, banking regulations, asset allocation, and risk management.

First, the findings show that the BCCH's actions on its policy rate and other monetary policy measures can have a noticeable effect on the interbank swap yield through the changes in the short-term interest rate. An increase (decrease) in the short-term interest rate is associated with

the rise (fall) in the swap yield. This shows that the BCCH's monetary policy can have a powerful impact on the financial conditions of financial institutions and other enterprises, as well as on the financial system.

Second, the empirical findings show that a rise (decline) in inflation and the growth of industrial production is usually associated with a higher (lower) swap yield, even though these effects are not statistically significant. This means that when the effective demand increases (decreases), which is often associated with inflation and the growth of industrial production, the interbank swap yield will tend to rise (fall).

Third, the findings associate the rise (decline) of the log of the equity price index with a higher (lower) swap yield. This implies that the swap buyer will receive a higher (lower) variable interest payments on swaps when financial markets are buoyant (sluggish) and rising (declining).

Fourth, the findings also relate the depreciation (appreciation) of the log of the exchange rate (the Chilean peso per US dollar) to a higher (lower) swap yield. This means that as the Chilean peso depreciates (appreciates) the buyer of the interest rate swap will receive a higher (lower) variable interest payment from the seller of the interest rate swap.

The empirical findings of the paper reinforce the view that the central bank can exert enormous influence on financial markets. The findings also support the notion that the central bank's actions influence the pricing of fixed income securities and derivatives, such as interbank interest rate swaps, through its monetary policy (Bindseil 2004; Fullwiler [2008] 2017).

VIII. CONCLUSION

The empirical analysis reveals that the month-over-month change in the short-term interest rate has an economically and statistically significant effect on the month-over-month change in the interbank swap yield of different maturity tenors after controlling for the month-over-month changes in several important macroeconomic and financial variables, such as inflation, the

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growth of industrial production, the logarithm of the equity price index, and the logarithm of the exchange rate. This shows that the central bank's monetary policy action, through its effect on the change in the short-term interest rate, influences the interbank swap yield. This finding demonstrates the central bank's ability to influence financial institutions' borrowing and lending rates over different time horizons. Given the growing importance of interbank interest rate swaps and other derivatives on the banking industry, financial intermediation, financial markets, and corporate finance, it is a substantive finding with implications for monetary policy, bank regulations, asset allocation, and risk management. This finding has repercussions for the private sector's marginal efficiency of capital, investment decisions, profitability, and leveraging decisions.

There is a paucity of literature on the empirical modeling of the interbank swap yield. The findings of this paper illuminate the macroeconomic and financial factors that produce interbank swap yield dynamics. These findings are not only germane to understanding such dynamics in Chile and other Latin American countries, but also elsewhere in both emerging markets and advanced countries. It is hoped that these findings with generate more detailed empirical studies of swap yields in other emerging markets and advanced countries and advance the empirical modeling of long-term swap yields.

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APPENDIX A: CORRELOGRAMS FOR GARCH(1,1) MODELS WITH ∆PDBC90D

Autocorrelation	Partial Correlation		AC	PAC	Q-Stat	Prob*
ı	ı <u>b</u> ı	1	0.050	0.050	0.5081	
ı İ i	1 1	2	-0.014	-0.017	0.5485	0.459
ı D ı	i ı 🗖 ı	3	0.071	0.072	1.5782	0.454
ı İ ı	111	4	-0.010	-0.018	1.6011	0.659
ı İ ı	j i j i	5	-0.039	-0.036	1.9232	0.750
ı İ ı	1	6	-0.008	-0.010	1.9368	0.858
i 🖬 i	(0)	7	-0.063	-0.062	2.7716	0.837
1 1		8	-0.008	0.004	2.7846	0.904
i ∳ i		9	0.014	0.012	2.8236	0.945
1 1		10	0.002	0.008	2.8241	0.971
i 🖞 i	ן ון ו	11	-0.027	-0.029	2.9776	0.982
↓ .		12	0.009	0.005	2.9949	0.991
1 0 1	ן ון ו	13	-0.027	-0.030	3.1519	0.994
ı þ í	ı <u>þ</u> ı	14	0.039	0.044	3.4870	0.996
1 1		15	0.004	-0.003	3.4908	0.998
C, i		16	-0.115	-0.112	6.4210	0.972
C, i		17	-0.119	-0.116	9.5930	0.887
I 🗖 I	וםי	18	-0.067	-0.068	10.601	0.877
I 🖣 I	111	19	-0.044	-0.024	11.035	0.893
I I I	111	20	-0.023	-0.011	11.154	0.919
1 D 1	ן ו	21	0.060	0.066	11.968	0.917
L i	l 🗖 I	22	-0.143	-0.164	16.622	0.734
1 1	11	23	0.004	-0.002	16.627	0.784
I 🗖 I	l i ∎i	24	-0.068	-0.114	17.685	0.774
	111	25	-0.044	-0.021	18.129	0.797
· ا	l 1 🗖 1	26	0.089	0.092	19.969	0.748
i 🕻 i	וםי	27	-0.041	-0.060	20.368	0.774
I 🗖 I	1 1	28	-0.067	-0.066	21.417	0.766
111	1 1	29	0.014	-0.041	21.463	0.806
ı 🔁		30	0.120	0.131	24.888	0.684
· 🗗		31	0.104	0.120	27.484	0.598
1 1	1 1	32	0.002	-0.004	27.485	0.648
111	I I	33	0.017	-0.037	27.553	0.691
1 j 1		34	0.028	-0.026	27.742	0.726
· P·	I I	35	0.069	0.039	28.924	0.715
· 🗐 ·	l ı ⊨ ı	36	0.072	0.103	30.210	0.699

Table A1. \triangle SWAP2Y = $\phi^1(C, \triangle$ PDBC90D, AR(1))

Autocorrelation	Partial Correlation		AC	PAC	Q-Stat	Prob*
ı D ı	.0.	1	0.047	0.047	0.4557	
101	1 1	2	-0.026	-0.028	0.5955	0.440
ı D ı	j , <u>p</u> ,	3	0.084	0.087	2.0513	0.359
111	1 1	4	-0.022	-0.031	2.1502	0.542
10		5	-0.026	-0.018	2.2885	0.683
I I I		6	-0.009	-0.016	2.3067	0.805
ı 🖬 i	ן ום ו	7	-0.066	-0.062	3.2261	0.780
1 1		8	-0.002	0.007	3.2269	0.863
1 1		9	-0.003	-0.007	3.2294	0.919
1 1 1	ן וויי	10	0.024	0.035	3.3553	0.949
1 (1		11	-0.038	-0.046	3.6701	0.961
1 j 1		12	0.010	0.015	3.6923	0.978
111		13	-0.012	-0.023	3.7234	0.988
ı þ í	ן וויי	14	0.033	0.040	3.9563	0.992
1 1		15	0.005	-0.003	3.9609	0.996
I <mark>L</mark> I		16	-0.105	-0.104	6.4171	0.972
E I		17	-0.129	-0.125	10.151	0.859
I <mark>I</mark> I	וםי	18	-0.066	-0.069	11.131	0.850
1 1 1	111	19	-0.040	-0.021	11.484	0.873
111		20	-0.019	-0.007	11.568	0.903
ı þ i	ı <u>þ</u> ı	21	0.061	0.072	12.412	0.901
 – – – – – – – – – – – – – – – – – – –	🗖 -	22	-0.147	-0.171	17.337	0.690
1 1	1 1	23	0.002	0.004	17.339	0.744
I 🗖 I	•	24	-0.064	-0.121	18.279	0.742
1 (1	111	25	-0.044	-0.013	18.740	0.766
ı þ i	ļ i 🏼 i	26	0.089	0.086	20.606	0.714
1 (1	101	27	-0.035	-0.045	20.899	0.747
I L I	l 1 0 1	28	-0.084	-0.084	22.557	0.709
ו 🎙 ו	111	29	0.032	-0.022	22.794	0.743
· 🗐 ·	ı 	30	0.101	0.117	25.215	0.667
۱ <mark>ا</mark>	ı 	31	0.106	0.116	27.920	0.575
1 1	1 1	32	-0.003	0.006	27.922	0.625
I I I	I 	33	0.016	-0.046	27.983	0.670
ı þ í	I []	34	0.027	-0.025	28.157	0.707
ı þ i	ן ו	35	0.064	0.034	29.182	0.703
i þi	<u> </u>	36	0.066	0.094	30.271	0.696

Table A2. Δ SWAP2Y = $\varphi^2(C, \Delta$ PDBC90D, Δ CPIYOY, Δ IPYOY, AR(1))

-

Autocorrelation	Partial Correlation		AC	PAC	Q-Stat	Prob*
	.]].	1	0.059	0.059	0.7216	
10	101	2	-0.040	-0.043	1.0479	0.306
ı D ı	ı 🗖 i	3	0.086	0.091	2.5685	0.277
ı İ ı	101	4	-0.021	-0.034	2.6598	0.447
ı 🖞 ı	101	5	-0.047	-0.036	3.1195	0.538
1 1		6	0.005	0.001	3.1255	0.681
I 🖬 I	10	7	-0.069	-0.070	4.1449	0.657
I 🗍 I	11	8	-0.008	0.008	4.1596	0.761
1 1	111	9	-0.007	-0.016	4.1698	0.841
1 1 1	I]I	10	0.018	0.031	4.2403	0.895
I 🗖 I	וםי	11	-0.071	-0.080	5.3318	0.868
1 🕴 1	I I	12	0.022	0.031	5.4340	0.908
1 🕴 1		13	-0.011	-0.027	5.4614	0.941
1 1	I]I	14	0.016	0.031	5.5203	0.962
1 D 1	111	15	0.030	0.019	5.7138	0.973
ı d i i		16	-0.113	-0.122	8.5466	0.900
C, I	l I C I	17	-0.118	-0.100	11.644	0.768
I 🚺 I	וםי	18	-0.040	-0.056	11.997	0.800
I 🚺 I	111	19	-0.040	-0.016	12.361	0.828
10	111	20	-0.031	-0.023	12.577	0.860
· ا	ן ו	21	0.082	0.089	14.119	0.824
C		22	-0.133	-0.168	18.191	0.637
1 1	111	23	0.002	0.021	18.192	0.695
I 🗖 I		24	-0.086	-0.156	19.909	0.647
1 ()	1 1	25	-0.045	-0.004	20.374	0.675
1 D 1	ן ו	26	0.081	0.077	21.908	0.641
111	10	27	-0.011	-0.041	21.934	0.692
I 🗖 I	101	28	-0.079	-0.070	23.396	0.664
I 🛛 I	10	29	0.029	-0.029	23.601	0.702
· P	l • 🗖	30	0.122	0.140	27.191	0.561
· 🗖 ·	l I 🗖 I	31	0.106	0.094	29.917	0.470
I 🕴 I		32	-0.017	0.001	29.985	0.518
1 j 1	I []	33	0.038	-0.038	30.332	0.551
I 🖡 I	I I I I	34	-0.011	-0.042	30.362	0.599
ı p ı	ן ו	35	0.067	0.050	31.485	0.592
ı 🗖 i		36	0.101	0.110	34.013	0.516

Table A3. Δ SWAP2Y = $\phi^{3}(C, \Delta$ PDBC90D, Δ CPIYOY, Δ IPYOY, Δ LNIGPA, Δ LNCLP, AR(1))

Autocorrelation	Partial Correlation		AC	PAC	Q-Stat	Prob*
. b i	l (b)	1	0.091	0.091	1.7133	
		2	0.035	0.027	1.9643	0.161
ı D ı	ı <u>D</u> ı	3	0.055	0.050	2.5992	0.273
ı d i	ן וער	4	-0.066	-0.077	3.5013	0.321
ı 🖞 i	ן ון ו	5	-0.058	-0.049	4.2078	0.379
ı İ ı	j i <u>j</u> i	6	0.027	0.039	4.3646	0.498
ı 🗖 i	j indi	7	-0.091	-0.087	6.1002	0.412
ı 🖞 ı	ן ון ו	8	-0.063	-0.049	6.9503	0.434
I I I		9	-0.012	-0.008	6.9816	0.539
ı 🖞 i	ן ון ו	10	-0.046	-0.031	7.4440	0.591
1 1		11	0.002	0.007	7.4447	0.683
ı İp i	ן וףי	12	0.056	0.041	8.1170	0.703
1 1		13	0.007	0.002	8.1291	0.775
ı þ i		14	0.031	0.018	8.3348	0.821
ı D i	ן וףי	15	0.080	0.061	9.7593	0.780
ı d ı	[]	16	-0.114	-0.128	12.663	0.628
i i i i i i i i i i i i i i i i i i i	•	17	-0.122	-0.115	15.989	0.454
1 1		18	0.004	0.023	15.993	0.524
ı d ı	ן ום י	19	-0.092	-0.060	17.883	0.463
1 🚺 1		20	-0.008	0.009	17.898	0.529
ı þ i	ן וףי	21	0.046	0.026	18.378	0.562
i i i i i i i i i i i i i i i i i i i	[]	22	-0.128	-0.119	22.104	0.394
ı þ í	ן וףי	23	0.026	0.037	22.255	0.445
1 0 1	10 1	24	-0.047	-0.091	22.768	0.474
I 🗖 I	10 1	25	-0.104	-0.094	25.297	0.390
ı p ı	ן וףי	26	0.069	0.060	26.423	0.385
I 🚺 I	10 1	27	-0.046	-0.089	26.929	0.413
∎, i	[]	28	-0.137	-0.120	31.381	0.256
ı p i	ן וףי	29	0.052	0.049	32.020	0.274
· ⊨	ļ ı 📮	30	0.159	0.172	38.079	0.121
I I I		31	-0.008	-0.012	38.093	0.147
ı p ı		32	0.052	-0.011	38.747	0.160
1 j 1	ן ון	33	0.017	-0.036	38.815	0.189
ı þ i	ן וים	34	0.038	0.068	39.162	0.213
ı p ı	ļ ī þ ī	35	0.068	0.036	40.313	0.211
i		36	0.059	0.015	41.168	0.219

Table A4. \triangle SWAP5Y = φ^4 (C, \triangle PDBC90D, AR(1))

Autocorrelation	Partial Correlation		AC	PAC	Q-Stat	Prob*
ı b ı	l 1 b i	1	0.086	0.086	1.5305	
		2	0.021	0.013	1.6176	0.203
ı D ı	i .	3	0.085	0.083	3.1069	0.212
101		4	-0.081	-0.096	4.4565	0.216
ı 🚺 i	j i j i	5	-0.048	-0.036	4.9409	0.293
ı İ I	i) i	6	0.024	0.027	5.0597	0.409
I 🗖 I	I []I	7	-0.093	-0.084	6.9028	0.330
10		8	-0.049	-0.035	7.4118	0.387
10	(9	-0.032	-0.035	7.6281	0.471
10	111	10	-0.040	-0.016	7.9656	0.538
1 1		11	0.000	0.001	7.9657	0.632
ı þ i	ן ו	12	0.054	0.046	8.5973	0.659
1 1 1		13	0.024	0.016	8.7199	0.727
ı þ í		14	0.038	0.022	9.0376	0.770
ı þ i	ן ו	15	0.063	0.044	9.9057	0.769
ı d ı		16	-0.099	-0.114	12.099	0.672
e i		17	-0.130	-0.124	15.868	0.462
1 1		18	0.016	0.033	15.925	0.529
I 🗖 I	וםי	19	-0.090	-0.061	17.732	0.473
111		20	-0.021	0.003	17.829	0.534
ı þ i		21	0.043	0.024	18.250	0.571
E I	I [] I	22	-0.118	-0.102	21.433	0.433
1 1	l i i i	23	0.012	0.021	21.467	0.492
I 🚺 I	I []I	24	-0.048	-0.093	21.997	0.520
I <mark>I</mark> II	I []I	25	-0.109	-0.093	24.741	0.420
ı p ı	ן ו	26	0.071	0.050	25.920	0.412
I 🚺 I	ן ום ו	27	-0.040	-0.072	26.295	0.447
 – – – – – – – – – – – – – – – – – – –	[]	28	-0.153	-0.143	31.852	0.238
ı p ı	ן ו	29	0.066	0.061	32.904	0.239
۱ ۵	ļ ı 🗖	30	0.137	0.161	37.415	0.136
111	111	31	-0.015	-0.018	37.468	0.164
ו 🎙 ו	1 1	32	0.055	-0.005	38.214	0.174
1 1	I []	33	0.024	-0.031	38.357	0.203
ı þ í	ן ו	34	0.028	0.052	38.552	0.233
ı p ı	ן ו	35	0.065	0.036	39.609	0.234
i þi	I]I	36	0.059	0.025	40.470	0.242

Table A5. Δ SWAP5Y = $\varphi^5(C, \Delta$ PDBC90D, Δ CPIYOY, Δ IPYOY, AR(1))

Autocorrelation	Partial Correlation		AC	PAC	Q-Stat	Prob*
ı 🗖 i	ı <u>b</u> ı	1	0.090	0.090	1.6652	
1 1	1 1	2	0.000	-0.008	1.6652	0.197
ı İ Di	j i pi	3	0.086	0.087	3.1959	0.202
ı 🗋 i		4	-0.076	-0.093	4.3825	0.223
ı 🖞 ı	ן ון ו	5	-0.045	-0.029	4.8094	0.307
1 j 1	ן וףי	6	0.036	0.035	5.0825	0.406
I 🗖 I	10 1	7	-0.075	-0.069	6.2743	0.393
1 ()		8	-0.036	-0.022	6.5447	0.478
1 ()	ן ון ו	9	-0.039	-0.049	6.8762	0.550
1 1 1		10	-0.038	-0.014	7.1911	0.617
1 1 1	ון ו	11	-0.041	-0.042	7.5623	0.672
ı þ í	ן וףי	12	0.031	0.036	7.7751	0.733
ı D ı	I I	13	0.030	0.024	7.9651	0.788
1 1	1 1	14	0.015	0.008	8.0173	0.842
i 🗖 i	ן ו	15	0.070	0.057	9.1086	0.824
I 🗖 I	•	16	-0.106	-0.130	11.582	0.710
∎, i	l III I	17	-0.124	-0.103	15.029	0.523
111		18	0.020	0.022	15.118	0.587
I 🗖 I	וםי	19	-0.077	-0.059	16.460	0.560
1 1		20	-0.003	0.019	16.462	0.626
1 D 1	1 1 1	21	0.049	0.017	17.012	0.652
I 🔲 I	וםי	22	-0.094	-0.079	19.055	0.582
1 1	1 1	23	0.002	0.011	19.056	0.642
I 🛛 I	l III I	24	-0.050	-0.084	19.637	0.664
ا ا	III	25	-0.113	-0.094	22.601	0.543
I D I	ן י ן י	26	0.058	0.048	23.394	0.555
		27	-0.009	-0.043	23.414	0.609
		28	-0.145	-0.148	28.395	0.391
		29	0.070	0.076	29.563	0.384
		30	0.132	0.135	33.740	0.249
1		31	-0.018	-0.017	33.822	0.288
1		32	0.024	-0.021	33.959	0.327
1 🛛 1		33	0.045	-0.017	34.460	0.351
1 [1		34	0.023	0.044	34.589	0.392
		35	0.064	0.038	35.612	0.392
I 🗐 I	I 	36	0.088	0.047	37.550	0.353

Table A6. \triangle SWAP5Y = $\varphi^{6}(C, \triangle$ PDBC90D, \triangle CPIYOY, \triangle IPYOY, \triangle LNIGPA, \triangle LNCLP, AR(1))

Autocorrelation	Partial Correlation		AC	PAC	Q-Stat	Prob*
, h u	l i h i	1	0.087	0.087	1.5388	
		2	0.008	0.000	1.5511	0.213
111	11	3	0.017	0.016	1.6108	0.447
ı 🗹 i	101	4	-0.072	-0.075	2.6895	0.442
ı 🖸 i	101	5	-0.062	-0.050	3.5057	0.477
ı İ ı	11	6	0.001	0.010	3.5057	0.623
ı 🗐 i	I I I	7	-0.096	-0.095	5.4510	0.487
ı 🗋 i	וםי	8	-0.072	-0.060	6.5511	0.477
111	111	9	-0.012	-0.010	6.5840	0.582
ı 🖸 i	וםי	10	-0.056	-0.055	7.2602	0.610
ı p ı	ן ו	11	0.064	0.064	8.1346	0.616
ı þ í	1 1	12	0.035	0.005	8.4065	0.676
1 1	1 1	13	0.006	-0.003	8.4148	0.752
11	101	14	-0.026	-0.046	8.5658	0.805
r 🆻 i	l i 🗖 i	15	0.084	0.083	10.122	0.753
10	101	16	-0.027	-0.037	10.283	0.802
I 🗖 I	ı ت ا	17	-0.078	-0.086	11.656	0.767
111	1 1	18	0.009	0.021	11.676	0.819
יםי	101	19	-0.072	-0.059	12.859	0.800
111	1 1	20	-0.023	-0.004	12.978	0.840
111	1 1	21	0.023	0.010	13.094	0.873
E, i		22	-0.137	-0.146	17.420	0.685
1 D 1	 	23	0.095	0.125	19.479	0.616
101	l III I	24	-0.039	-0.100	19.840	0.652
Q I		25	-0.130	-0.115	23.770	0.475
ı p ı	ן ו	26	0.056	0.046	24.504	0.490
I 🛛 I	l III I	27	-0.062	-0.110	25.409	0.496
יםי	101	28	-0.080	-0.047	26.934	0.467
ı D ı	1 1	29	0.053	0.020	27.598	0.486
· 🗖	l I 🗖 I	30	0.109	0.095	30.425	0.393
I D I	111	31	0.027	0.011	30.603	0.435
ı D ı		32	0.071	0.018	31.821	0.425
I 🛛 I	101	33	-0.053	-0.059	32.504	0.442
ı D ı	I <mark>]</mark>	34	0.032	0.037	32.754	0.479
ı p ı	I <mark>]</mark>	35	0.048	0.039	33.333	0.500
· ۲	<u> </u>	36	0.062	0.050	34.283	0.503

Table A7. \triangle SWAP10Y = ϕ^7 (C, \triangle PDBC90D, AR(1))

Autocorrelation	Partial Correlation		AC	PAC	Q-Stat	Prob*
. Di	l (b)	1	0.088	0.088	1.5812	
ı [ı		2	-0.002	-0.010	1.5821	0.208
1 🛛 1	1	3	0.040	0.041	1.9074	0.385
ı d i	i idii	4	-0.089	-0.097	3.5497	0.314
ı 🚺 i	i i i i	5	-0.046	-0.029	3.9946	0.407
1 1	1 1	6	-0.007	-0.004	4.0053	0.549
ı 🗖 i	i _	7	-0.097	-0.090	5.9826	0.425
ı 🗋 i	ום ו	8	-0.061	-0.051	6.7825	0.452
I 🚺 I		9	-0.033	-0.033	7.0189	0.535
I 🗍 I		10	-0.041	-0.032	7.3795	0.598
ı þ i	ı <u>þ</u> ı	11	0.063	0.058	8.2351	0.606
ı þ í		12	0.033	0.008	8.4781	0.670
1 1 1		13	0.018	0.009	8.5498	0.741
11	(]	14	-0.023	-0.048	8.6621	0.798
ı D ı	ן וים	15	0.069	0.075	9.7227	0.782
	10 1	16	-0.010	-0.026	9.7451	0.835
I 🗖 I	I [] I	17	-0.092	-0.095	11.640	0.768
1	ן ו	18	0.024	0.036	11.765	0.814
I 🗖 I	וםי	19	-0.077	-0.069	13.091	0.786
1 1 1	1 1	20	-0.029	0.001	13.277	0.824
1 1	1 1	21	0.024	0.006	13.404	0.859
Eļ i	•	22	-0.134	-0.134	17.500	0.680
i 🗖 i	ļ i 🎫	23	0.080	0.106	18.982	0.646
11	I I ⊑ I	24	-0.041	-0.099	19.368	0.680
E, i	I II I	25	-0.130	-0.107	23.288	0.503
ı 🏼 i	ן ו	26	0.055	0.032	24.002	0.519
I 🛛 I		27	-0.055	-0.090	24.725	0.535
I 🔲 I	10 1	28	-0.091	-0.066	26.688	0.481
I 🛛 I	1 1	29	0.060	0.024	27.537	0.489
1 D 1	ן ו	30	0.093	0.092	29.624	0.433
111	1 1	31	0.023	-0.006	29.755	0.478
1 D 1	ן ו	32	0.069	0.033	30.916	0.470
10	I [] I	33	-0.047	-0.063	31.465	0.494
111	1 I I I	34	0.023	0.024	31.599	0.537
ı p ı	l i 🏼 i	35	0.048	0.046	32.175	0.557
· ۱	ן ו	36	0.061	0.048	33.094	0.560

Table A8. \triangle SWAP10Y = $\varphi^{8}(C, \triangle$ PDBC90D, \triangle CPIYOY, \triangle IPYOY, AR(1))

Autocorrelation	Partial Correlation		AC	PAC	Q-Stat	Prob*
. Di	.	1	0.092	0.092	1.7420	
ı İ ı	ן ון ו	2	-0.020	-0.029	1.8226	0.177
ı D ı	ן ואַ ד	3	0.063	0.068	2.6464	0.266
ı d i		4	-0.085	-0.100	4.1543	0.245
ı 🗍 i		5	-0.027	-0.005	4.3069	0.366
1 1		6	-0.005	-0.012	4.3124	0.505
1 0 1	ן ום ו	7	-0.066	-0.054	5.2341	0.514
ı () i	ן ון ו	8	-0.038	-0.033	5.5357	0.595
ı () i	(9	-0.031	-0.031	5.7404	0.676
ı 🖣 i	(10	-0.038	-0.028	6.0538	0.735
ı j i		11	0.025	0.024	6.1871	0.799
1 1		12	-0.008	-0.020	6.2011	0.860
ı j ı	ן וויי	13	0.035	0.038	6.4653	0.891
1 ()	ן וםי	14	-0.037	-0.061	6.7697	0.914
ı D ı	ן וים	15	0.047	0.063	7.2633	0.924
111	ן ון ו	16	-0.014	-0.042	7.3088	0.949
ال اً ا	ים,	17	-0.103	-0.090	9.6867	0.882
1 1		18	-0.003	-0.000	9.6885	0.916
I 🚺 I	ן וםי	19	-0.053	-0.053	10.332	0.921
I I I		20	-0.014	0.011	10.377	0.943
1] 1		21	0.027	0.003	10.543	0.957
	ן מי	22	-0.113	-0.116	13.462	0.892
I PI	ļ ' P '	23	0.067	0.091	14.510	0.882
I 🛛 I		24	-0.040	-0.094	14.874	0.899
٩		25	-0.135	-0.103	19.119	0.746
I DI	I]I	26	0.059	0.035	19.925	0.751
		27	-0.031	-0.047	20.146	0.785
		28	-0.096	-0.086	22.339	0.720
		29	0.062	0.041	23.267	0.720
		30	0.056	0.051	24.024	0.728
		31	0.028	0.011	24.212	0.762
I ∐ I		32	0.058	0.024	25.024	0.767
		33	-0.041	-0.058	25.435	0.788
I ∐ I . 6		34	0.048	0.046	25.999	0.802
1 H 1			0.048	0.034	20.570	0.814
· •		36	0.038	0.019	26.927	0.834

Table A9. Δ SWAP10Y = $\phi^9(C, \Delta$ PDBC90D, Δ CPIYOY, Δ IPYOY, Δ LNIGPA, Δ LNCLP, AR(1))

APPENDIX B: CORRELOGRAMS FOR GARCH(1,1) MODELS WITH ΔPDBC30D

Autocorrelation	Partial Correlation		AC	PAC	Q-Stat	Prob*
. h .	h .	1	0.035	0.035	0 2502	
· · ·		2	0.055	0.054	0.8826	0.347
· p·	,	3	0.052	0.049	1.4453	0.485
		4	0.004	-0.002	1,4489	0.694
		5	-0.070	-0.076	2.4625	0.651
1 🛛 1		6	-0.012	-0.010	2.4922	0.778
1 1	1 1	7	-0.005	0.004	2.4974	0.869
101	101	8	-0.065	-0.056	3.3826	0.848
ı 🛛 ı	ı 🗓 ı	9	0.032	0.038	3.5999	0.891
1 1	111	10	-0.005	-0.006	3.6060	0.935
10	I ()	11	-0.046	-0.046	4.0665	0.944
ı 🗖 i	l i þ i	12	0.085	0.087	5.6329	0.897
10	10	13	-0.058	-0.068	6.3559	0.897
ı 🛛 ı	I 🗍 I	14	0.044	0.049	6.7779	0.913
ı ()	ן ומי	15	-0.038	-0.044	7.1029	0.931
ı d ı		16	-0.091	-0.101	8.9180	0.882
		17	-0.157	-0.138	14.424	0.567
I 🖸 I	ום	18	-0.072	-0.062	15.592	0.553
ı Q ı	ון	19	-0.062	-0.036	16.445	0.562
ום	111	20	-0.055	-0.023	17.129	0.581
ı D ı	l I 🗖 I	21	0.112	0.105	19.960	0.460
I <mark>E</mark> I		22	-0.114	-0.134	22.955	0.346
1	111	23	-0.010	-0.022	22.978	0.403
I 🗖 I	l I E I	24	-0.068	-0.106	24.043	0.401
I (I)	l (l 1	25	-0.047	-0.043	24.561	0.430
ı 🛛 ı	l I 🗖 I	26	0.056	0.075	25.290	0.446
I C I	וםי	27	-0.078	-0.098	26.723	0.424
I C I	l I E I	28	-0.078	-0.096	28.146	0.403
ı D ı	ון ו	29	0.032	0.051	28.394	0.444
۱ þ	l I 🗖 I	30	0.128	0.114	32.323	0.306
ı 🛛 ı	· ا	31	0.061	0.090	33.218	0.313
ו 🏾 ד	l I 🛛 I	32	0.063	0.033	34.175	0.318
ı D ı	וםיו	33	0.045	-0.066	34.674	0.342
1 1	I I I I	34	0.022	0.016	34.792	0.383
ı þ í	I ()	35	0.031	-0.032	35.034	0.419
ı þ i		36	0.097	0.123	37.385	0.360

Table B1. \triangle SWAP2Y = $\psi^1(C, \triangle$ PDBC30D, AR(1))

Autocorrelation	Partial Correlation		AC	PAC	Q-Stat	Prob*
. b .	ı <u>b</u> ı	1	0.031	0.031	0.1970	
ı D ı	i 101	2	0.046	0.045	0.6307	0.427
ı D ı	ן ו	3	0.065	0.062	1.5014	0.472
1 1	111	4	-0.006	-0.012	1.5098	0.680
ı 🖬 i	101	5	-0.060	-0.066	2.2739	0.686
111	111	6	-0.011	-0.011	2.3013	0.806
1 1	1 1	7	-0.008	-0.000	2.3143	0.889
I 🛛 I	10	8	-0.054	-0.045	2.9399	0.891
1 1 1	111	9	0.016	0.019	2.9913	0.935
1 1 1		10	0.012	0.012	3.0225	0.963
1 0 1	ן ום י	11	-0.056	-0.054	3.6925	0.960
ı þ i	ı þ i	12	0.088	0.087	5.3546	0.913
1 0 1	ן ום י	13	-0.041	-0.049	5.7150	0.930
ı þ í	ן ו	14	0.039	0.043	6.0433	0.945
1 1 1	ום ו	15	-0.042	-0.052	6.4327	0.954
I 🗖 I	ı []ı	16	-0.083	-0.088	7.9758	0.925
 •		17	-0.159	-0.149	13.619	0.627
ı Q ı	ום ו	18	-0.071	-0.054	14.736	0.614
I 🗖 I	l I C I	19	-0.061	-0.040	15.583	0.622
ı ğ ı	I (I	20	-0.054	-0.025	16.255	0.640
ı p ı	l i 🗖 i	21	0.112	0.115	19.127	0.514
I L I I		22	-0.114	-0.140	22.123	0.392
111	111	23	-0.011	-0.014	22.149	0.451
I 🗖 I	l I C I	24	-0.065	-0.111	23.117	0.454
I 🗖 I	l (l 1	25	-0.056	-0.046	23.847	0.470
ı p ı	ן ו	26	0.060	0.071	24.689	0.480
I 🗖 I	l (E)	27	-0.072	-0.082	25.898	0.469
I L I	l I C I	28	-0.092	-0.113	27.916	0.415
ו 🎙 ו	ן ו	29	0.040	0.059	28.302	0.449
ı ⊨ ı	l I 🗐 I	30	0.113	0.111	31.361	0.349
ı pı	l i þ í	31	0.069	0.091	32.504	0.344
ı p ı	ן ו	32	0.055	0.038	33.241	0.359
ı p ı	I I D I	33	0.046	-0.071	33.757	0.383
111		34	0.022	0.014	33.880	0.425
ı þ i	I I I I	35	0.025	-0.034	34.039	0.466
<u> </u>		36	0.094	0.117	36.254	0.410

Table B2. \triangle SWAP2Y = $\psi^2(C, \triangle$ PDBC30D, \triangle CPIYOY, \triangle IPYOY, AR(1))

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Autocorrelation	Partial Correlation		AC	PAC	Q-Stat	Prob*
. D .	101	1	0.039	0.039	0.3167	
, D (2	0.026	0.025	0.4580	0.499
ı D ı	i i <u>b</u> i	3	0.061	0.059	1.2235	0.542
1 1	i i i i	4	0.001	-0.004	1.2236	0.747
ı d ı	ן ום ו	5	-0.066	-0.069	2.1342	0.711
1 1		6	0.005	0.006	2.1387	0.830
1 1		7	0.003	0.006	2.1404	0.906
ı 🖬 i	ן ומי	8	-0.066	-0.059	3.0582	0.880
ı İ ı		9	0.028	0.032	3.2239	0.920
1 1		10	0.009	0.005	3.2410	0.954
ı d i	101	11	-0.087	-0.083	4.8841	0.899
ı þ i		12	0.085	0.091	6.4580	0.841
1 (1	(13	-0.035	-0.048	6.7283	0.875
1 1 1	ן וויין	14	0.015	0.029	6.7786	0.913
1 (1	10 1	15	-0.025	-0.033	6.9167	0.938
I L I	l I C I	16	-0.093	-0.106	8.8139	0.887
 – – – – – – – – – – – – – – – – – – –		17	-0.147	-0.128	13.644	0.625
I 🗖 I	(18	-0.059	-0.047	14.428	0.637
101	l (l 1	19	-0.049	-0.040	14.962	0.665
I 🚺 I	111	20	-0.056	-0.024	15.678	0.679
, ⊫		21	0.131	0.129	19.589	0.484
۱ Щ ۱	l Q '	22	-0.096	-0.129	21.695	0.417
111	1 1	23	-0.022	-0.007	21.808	0.471
1 [] 1	•	24	-0.077	-0.126	23.184	0.450
10	(])	25	-0.053	-0.047	23.848	0.470
ı Dı	ן ו	26	0.047	0.074	24.367	0.498
1 [] 1	ן יםי	27	-0.040	-0.065	24.752	0.533
I 🔲 I		28	-0.085	-0.103	26.460	0.493
1 D 1	ן ו	29	0.037	0.060	26.778	0.530
· P	ļ I P I	30	0.124	0.104	30.469	0.391
ı p ı	I I I I I I I I	31	0.066	0.078	31.509	0.391
1 D 1		32	0.037	0.036	31.843	0.424
ı p ı	I []	33	0.067	-0.051	32.946	0.421
1 1		34	0.004	0.005	32.950	0.470
ı D ı		35	0.033	-0.016	33.215	0.506
· 🍽		36	0.110	0.114	36.199	0.412

Table B3. \triangle SWAP2Y = $\psi^{3}(C, \triangle$ PDBC30D, \triangle CPIYOY, \triangle IPYOY, \triangle LNIGPA, \triangle LNCLP, AR(1))

Autocorrelation	Partial Correlation		AC	PAC	Q-Stat	Prob*
, h i	l i hni	1	0.082	0.082	1.3755	
, D ,	, , , , , , , , , , , , , , , , , , ,	2	0.039	0.032	1.6838	0.194
ı D ı	, , , , , , , , , , , , , , , , , , ,	3	0.045	0.040	2.1093	0.348
101		4	-0.078	-0.087	3.3695	0.338
I E I		5	-0.093	-0.084	5.1741	0.270
ı İ ı	ן וויין	6	0.031	0.050	5.3820	0.371
10	ן ומי	7	-0.067	-0.060	6.3174	0.389
ı [] I	10 1	8	-0.088	-0.081	7.9455	0.337
1 1		9	0.015	0.015	7.9929	0.434
ı ğ ı	I (I	10	-0.052	-0.045	8.5633	0.479
111	1 1	11	-0.013	-0.002	8.6005	0.570
ı p ı	ן ו	12	0.092	0.072	10.423	0.493
1 1	1 1	13	0.011	-0.004	10.451	0.576
ו 🎙 ו	ן ו	14	0.032	0.026	10.674	0.638
ו 🏾 ו	111	15	0.052	0.021	11.261	0.665
L		16	-0.120	-0.123	14.441	0.492
L	l I C I	17	-0.131	-0.110	18.247	0.310
1 1	111	18	-0.005	0.009	18.253	0.373
I L I	ן ום ו	19	-0.085	-0.052	19.865	0.340
1 1	111	20	0.002	0.020	19.866	0.403
ו 🏾 ד	ן ו	21	0.061	0.029	20.717	0.414
Q I		22	-0.115	-0.121	23.740	0.306
ו р ו	ן ו	23	0.044	0.057	24.177	0.338
I 🚺 I	l I C I	24	-0.046	-0.100	24.665	0.368
I L I	l 1 0 1	25	-0.112	-0.111	27.597	0.277
ו р ו	ן ו	26	0.045	0.044	28.062	0.305
1 0 1	l I C I	27	-0.058	-0.113	28.856	0.318
 •		28	-0.153	-0.128	34.384	0.155
ו р ו	ן ו	29	0.055	0.070	35.118	0.166
۱ 🗖 ۱	ļ ı 🗖	30	0.164	0.162	41.543	0.062
I 🚺 I	111	31	-0.027	-0.020	41.718	0.076
ı p ı	1 1	32	0.075	-0.001	43.079	0.073
ı þ í	I 	33	0.031	-0.044	43.319	0.087
1 1	ן וףי	34	0.014	0.052	43.369	0.107
ı þ i		35	0.042	-0.001	43.800	0.121
ı þ i	ום	36	0.084	0.046	45.548	0.109

Table B4. \triangle SWAP5Y = ψ^4 (C, \triangle PDBC30D, AR(1))

Autocorrelation	Partial Correlation		AC	PAC	Q-Stat	Prob*
. Di	l 1 0 1	1	0.074	0.074	1.1113	
ı D i	I <mark>]</mark>	2	0.035	0.029	1.3580	0.244
ı D ı	j i <u>p</u> i	3	0.076	0.071	2.5417	0.281
ı d i i	j jeji	4	-0.083	-0.096	3.9842	0.263
ı d i i	i idii	5	-0.089	-0.082	5.6522	0.227
ı İ ı	ı <u>þ</u> ı	6	0.030	0.043	5.8383	0.322
ı 🖬 i	ן ון ו	7	-0.061	-0.048	6.6194	0.357
ı 🗋 i	וםי	8	-0.067	-0.058	7.5667	0.372
I 🖡 I		9	-0.011	-0.019	7.5910	0.474
1 0 1	ן ון ו	10	-0.054	-0.042	8.2220	0.512
		11	-0.011	0.004	8.2464	0.605
ı 🗖 i	ן וףי	12	0.092	0.080	10.073	0.524
ı þ í		13	0.027	0.014	10.231	0.596
ı p ı	I I	14	0.038	0.022	10.553	0.648
ı p ı	1 1	15	0.032	0.001	10.783	0.703
I L I I		16	-0.105	-0.104	13.218	0.585
 •	•	17	-0.140	-0.127	17.597	0.348
1 1 1	ן וףי	18	0.010	0.029	17.619	0.413
I 🗖 I	ן ום ו	19	-0.089	-0.059	19.408	0.367
1		20	-0.020	0.000	19.495	0.426
ı þ i		21	0.049	0.024	20.046	0.455
ı d i i		22	-0.106	-0.104	22.602	0.366
ı þ í	ן וףי	23	0.030	0.044	22.816	0.412
1 1 1	III I	24	-0.049	-0.099	23.373	0.439
E I	[]	25	-0.121	-0.121	26.765	0.316
ı D ı	ן וףי	26	0.048	0.034	27.304	0.341
I 🖸 I	10 1	27	-0.052	-0.092	27.940	0.361
 •	🗖 -	28	-0.163	-0.147	34.251	0.159
ı pı	ן ו	29	0.062	0.070	35.180	0.165
· 🗖		30	0.136	0.150	39.627	0.090
1 1 1	1 1	31	-0.034	-0.023	39.899	0.107
ı pı	1 1	32	0.071	-0.003	41.132	0.105
ı þ i		33	0.042	-0.034	41.566	0.120
1 1	ן וףי	34	0.005	0.029	41.572	0.145
ı þ í		35	0.038	0.005	41.921	0.165
<u> </u>	ıþı	36	0.087	0.052	43.801	0.146

Table B5. Δ SWAP5Y = $\psi^{5}(C, \Delta$ PDBC30D, Δ CPIYOY, Δ IPYOY, AR(1))

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Autocorrelation	Partial Correlation		AC	PAC	Q-Stat	Prob*
. b i	l . Di	1	0.075	0.075	1.1532	
		2	0.017	0.012	1.2144	0.270
ı D ı		3	0.080	0.079	2.5534	0.279
ı 🗖 ı		4	-0.081	-0.094	3.9295	0.269
ı d i	101	5	-0.077	-0.066	5.1570	0.272
ı <mark>İ</mark> li	ן ו	6	0.041	0.049	5.5134	0.356
ı 🖞 i	111	7	-0.031	-0.022	5.7145	0.456
I 🗍 I	ום	8	-0.057	-0.052	6.4127	0.492
I I I	111	9	-0.010	-0.021	6.4352	0.599
I 🖬 I		10	-0.055	-0.045	7.0783	0.629
ı 🖬 i	10 1	11	-0.054	-0.037	7.7168	0.656
ı þ i	ן ו	12	0.053	0.051	8.3347	0.683
ı p ı	ן ו	13	0.038	0.033	8.6538	0.732
1	1 1	14	0.014	0.008	8.6988	0.795
ı p ı	1 1	15	0.034	0.007	8.9478	0.834
∎, i		16	-0.123	-0.134	12.325	0.654
Eļ i	I I E I	17	-0.127	-0.103	15.917	0.459
1 1	111	18	-0.003	0.010	15.920	0.530
I 🗖 I	1 1	19	-0.069	-0.049	16.981	0.524
1 1	111	20	0.005	0.016	16.986	0.591
1 D 1	ון ו	21	0.059	0.025	17.780	0.602
יםי	וםי	22	-0.074	-0.077	19.040	0.583
111	111	23	0.010	0.023	19.064	0.641
10	יםי	24	-0.049	-0.082	19.614	0.665
Q '	ן מי	25	-0.126	-0.120	23.302	0.502
I D I		26	0.035	0.029	23.585	0.543
		27	-0.020	-0.060	23.682	0.594
		28	-0.160	-0.159	29.745	0.326
		29	0.064	0.077	30.736	0.329
י 📮 י		30	0.120	0.126	34.191	0.232
۱ ۱ ۱		31	-0.034	-0.018	34.477	0.262
1 1		32	0.040	-0.022	34.860	0.289
1 1 1		33	0.065	-0.016	35.905	0.290
1 1 L		34	0.006	0.022	35.912	0.334
		35	0.038	0.008	36.267	0.363
· 🗖	ן ו	36	0.105	0.058	38.999	0.295

Table B6. \triangle SWAP5Y = $\psi^{6}(C, \triangle PDBC30D, \triangle CPIYOY, \triangle IPYOY, \triangle LNIGPA, \triangle LNCLP, AR(1))$

Autocorrelation	Partial Correlation		AC	PAC	Q-Stat	Prob*
ı b ı	l (b)	1	0.083	0.083	1.4189	
, j		2	0.010	0.003	1.4396	0.230
1 1	111	3	-0.003	-0.004	1.4411	0.486
10	101	4	-0.080	-0.080	2.7851	0.426
I 🗖 I	101	5	-0.083	-0.071	4.2240	0.377
1 1	1	6	-0.000	0.014	4.2240	0.518
ı d i	j i g i	7	-0.091	-0.093	5.9861	0.425
ı 🖬 i	I []	8	-0.073	-0.066	7.1138	0.417
1 1		9	0.006	0.006	7.1223	0.523
10	וםי	10	-0.050	-0.057	7.6509	0.570
ı þ i		11	0.046	0.042	8.1117	0.618
ı D i	ן ו	12	0.058	0.028	8.8500	0.636
1 1 1		13	0.010	-0.006	8.8710	0.714
111	ון ו	14	-0.019	-0.033	8.9475	0.777
ı þ i	ן ו	15	0.063	0.056	9.8260	0.775
1 1 1	I []	16	-0.031	-0.030	10.038	0.817
ı d ı	10	17	-0.089	-0.092	11.814	0.757
1 1	l I I I	18	0.011	0.022	11.839	0.810
ı Q ı	ן ום ו	19	-0.074	-0.060	13.058	0.788
111		20	-0.011	0.004	13.087	0.834
ı <mark>þ</mark> i	111	21	0.037	0.020	13.399	0.860
L		22	-0.134	-0.149	17.513	0.680
۱ <mark>ا</mark>	ı 	23	0.104	0.130	19.981	0.584
I 🚺 I	l 1 0 1	24	-0.036	-0.098	20.275	0.625
C		25	-0.133	-0.132	24.383	0.440
ו 🎙 ו	ן ו	26	0.026	0.033	24.546	0.488
I 🗖 I		27	-0.070	-0.124	25.689	0.480
I L I	ום ו	28	-0.089	-0.055	27.573	0.433
ו 🎙 ו	111	29	0.040	0.010	27.955	0.467
· 🖻	I 🖻 I	30	0.123	0.099	31.559	0.340
1 1	1 1	31	0.017	0.001	31.632	0.385
ı p ı	111	32	0.081	0.021	33.225	0.359
1 0 1	וםי	33	-0.056	-0.076	33.987	0.372
1 1	1 1 1	34	0.006	0.019	33.996	0.419
ı D ı	ן ו	35	0.045	0.034	34.491	0.444
<u> </u>	I I	36	0.083	0.065	36.200	0.412

Table B7. \triangle SWAP10Y = ψ^7 (C, \triangle PDBC30D, AR(1))

Autocorrelation	Partial Correlation		AC	PAC	Q-Stat	Prob*
. b i	l i h i	1	0.086	0.086	1.5034	
		2	0.005	-0.002	1.5089	0.219
111		3	0.022	0.022	1.6085	0.447
ı d i		4	-0.095	-0.099	3.4839	0.323
101	101	5	-0.072	-0.056	4.5615	0.335
111	1 1	6	-0.009	0.001	4.5788	0.469
ı d i	I E I	7	-0.089	-0.085	6.2381	0.397
ı 🗋 i	ום	8	-0.061	-0.054	7.0232	0.426
111	111	9	-0.015	-0.018	7.0704	0.529
1 1 1	10	10	-0.036	-0.035	7.3466	0.601
ı þ i	ı þ ı	11	0.045	0.038	7.7848	0.650
ı <mark>p</mark> ı	ן ו	12	0.058	0.031	8.5100	0.667
1 1 1	111	13	0.024	0.009	8.6381	0.733
111	I []	14	-0.013	-0.033	8.6735	0.797
ı þ i	ı <u>þ</u> ı	15	0.047	0.044	9.1565	0.821
111	111	16	-0.016	-0.017	9.2147	0.866
I L I I		17	-0.104	-0.103	11.628	0.769
ı þ í	ן ו	18	0.028	0.044	11.799	0.812
ı d ı	ן ום ו	19	-0.084	-0.079	13.406	0.767
111		20	-0.020	0.007	13.494	0.812
ı þ í	I I	21	0.034	0.016	13.762	0.842
d I		22	-0.131	-0.140	17.698	0.668
ı þ i	ı 1	23	0.089	0.111	19.539	0.612
1 (1	I []I	24	-0.041	-0.103	19.928	0.646
E I		25	-0.137	-0.127	24.300	0.445
1 1 1	111	26	0.024	0.016	24.440	0.494
ı 🖸 I	I []I	27	-0.063	-0.104	25.389	0.497
I <mark>n</mark> I	וםי	28	-0.099	-0.074	27.689	0.427
ı þ i	i i	29	0.043	0.009	28.131	0.458
ı 🗖 i	ı þ i	30	0.106	0.095	30.847	0.373
1 1 1	111	31	0.011	-0.019	30.878	0.421
ı þ i	ן ו	32	0.075	0.031	32.243	0.405
I 🚺 I	(1	33	-0.047	-0.078	32.779	0.429
1 1		34	-0.004	-0.001	32.784	0.478
ı þ i	ı <mark>þ</mark> i	35	0.046	0.047	33.297	0.502
. D i	l ı 🕅 ı	36	0.086	0.064	35.131	0.462

Table B8. \triangle SWAP10Y = ψ ⁸(C, \triangle PDBC30D, \triangle CPIYOY, \triangle IPYOY, AR(1))

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Autocorrelation	Partial Correlation		AC	PAC	Q-Stat	Prob*
ı b ı	ı b ı	1	0.087	0.087	1.5422	
111	111	2	-0.010	-0.018	1.5630	0.211
ı D ı	ן ו	3	0.045	0.048	1.9842	0.371
ı d i	j i g i	4	-0.092	-0.102	3.7640	0.288
ı 🗋 i		5	-0.052	-0.033	4.3254	0.364
1 1	1 1	6	-0.008	-0.006	4.3382	0.502
1 0 1	10	7	-0.053	-0.044	4.9221	0.554
1 (1	101	8	-0.038	-0.035	5.2317	0.632
1 1	111	9	-0.006	-0.009	5.2401	0.732
1 Q 1	I (I	10	-0.038	-0.037	5.5538	0.784
1 1	1 1	11	0.006	0.007	5.5619	0.851
1 1	1 1	12	0.006	-0.008	5.5686	0.901
ı p ı	ן ו	13	0.046	0.046	6.0270	0.915
10	101	14	-0.027	-0.047	6.1867	0.939
111	111	15	0.018	0.022	6.2571	0.960
101	1 1	16	-0.031	-0.044	6.4704	0.971
I <mark>I</mark> II	I I ⊑ I	17	-0.109	-0.096	9.1136	0.909
111	1 1	18	-0.011	-0.002	9.1411	0.936
101	וםי	19	-0.060	-0.061	9.9406	0.934
1 1	111	20	0.000	0.018	9.9406	0.954
ı p ı	111	21	0.043	0.018	10.366	0.961
I I I	ן מי	22	-0.109	-0.125	13.073	0.906
1 D 1		23	0.073	0.089	14.291	0.891
		24	-0.040	-0.090	14.661	0.906
	ן פיי	25	-0.148	-0.130	19.744	0.711
1] 1		26	0.031	0.019	19.965	0.749
		27	-0.037	-0.058	20.285	0.778
		28	-0.106	-0.098	22.925	0.689
1 1		29	0.036	0.011	23.228	0.722
1 H 1	· ·	30	0.064	0.055	24.216	0./18
I I I		31	0.021	0.004	24.325	0.757
I ∐ I		32	0.061	0.018	25.235	0.757
		33	-0.041	-0.077	25.639	0.779
		34	0.017	0.015	25./11	0.813
		35	0.047	0.034	26.250	0.826
۱ ۴ ۱	I 	36	0.057	0.028	27.069	0.829

Table B9. Δ SWAP10Y = ψ^9 (C, Δ PDBC30D, Δ CPIYOY, Δ IPYOY, Δ LNIGPA, Δ LNCLP, AR(1))

APPENDIX C: ADDITIONAL GARCH(1,1) MODELS

	∆SWAP2Y	∆SWAP2Y	∆SWAP5Y	∆SWAP5Y	∆SWAP10Y	∆SWAP10Y	
Mean Equation							
Intercept	-0.01	-0.01	-0.01	-0.02	-0.01	-0.02	
-	(0.77)	(0.60)	(0.59)	(0.54)	(0.62)	(0.36)	
∆PDBC90D	0.28	0.29	0.10	0.12	0.004	0.02	
	(0.00)	(0.00)	(0.02)	(0.00)	(0.94)	(0.75)	
<i>ACOREYOY</i>	-0.02	-0.02	-0.03	-0.06	-0.06	-0.08	
	(0.60)	(0.74)	(0.53)	(0.26)	(0.19)	(0.07)	
ΔΙΡΥΟΥ	0.003	0.002	0.004	0.002	0.003	0.002	
	(0.30)	(0.39)	(0.08)	(0.17)	(0.14)	(0.17)	
ΔLNIGPA		0.90		0.85		0.58	
		(0.01)		(0.03)		(0.18)	
ΔLNCLP		0.80		2.33		2.45	
		(0.39)		(0.00)		(0.00)	
AR(1)	0.38	0.36	0.29	0.29	0.27	0.27	
	(0.00)	(0.00)	(0.00)	(0.00)	(0.01)	(0.00)	
Variance Equation							
Intercept	0.01	0.01	0.004	0.004	0.004	0.003	
	(0.23)	(0.19)	(0.29)	(0.24)	(0.19)	(0.23)	
ARCH	0.15	0.19	0.11	0.15	0.13	0.16	
	(0.16)	(0.13)	(0.10)	(0.11)	(0.03)	(0.04)	
GARCH	0.59	0.56	0.79	0.78	0.77	0.78	
	(0.04)	(0.04)	(0.00)	(0.00)	(0.00)	(0.00)	
Model Information							
Obs	142	142	142	142	142	142	
Adj R ²	0.30	0.30	0.15	0.18	0.07	0.12	
AIC	- 0.23	-0.24	- 0.34	- 0.39	-0.47	- 0.53	
Diagnostic Tests							
ARCH LM	1.41	1.12	0.93	0.59	0.84	0.67	
(12 lags)	(0.17)	(0.35)	(0.52)	(0.84)	(0.61)	(0.78)	
DW Stat	1.96	1.93	1.87	1.79	1.88	1.80	
JQB	24.49	15.43	11.78	9.56	7.96	5.24	
	(0.00)	(0.00)	(0.00)	(0.01)	(0.02)	(0.07)	

Table C1. GARCH (1,1) Model (with \triangle PDBC90D and \triangle COREYOY)

Note: All vars are in diff, *p*-values are in parenthesis

	∆SWAP2Y	∆SWAP2Y	∆SWAP5Y	∆SWAP5Y	∆SWAP10Y	∆SWAP10Y		
Mean Equation								
Intercept	-0.01	-0.01	-0.01	-0.02	-0.01	-0.02		
	(0.83)	(0.58)	(0.64)	(0.34)	(0.63)	(0.64)		
∆PDBC30D	0.42	0.44	0.22	0.24	0.14	0.16		
	(0.00)	(0.00)	(0.00)	(0.00)	(0.02)	(0.00)		
<i>ACOREYOY</i>	-0.04	-0.03	-0.04	-0.06	-0.05	-0.07		
	(0.53)	(0.59)	(0.33)	(0.15)	(0.22)	(0.09)		
ΔΙΡΥΟΥ	0.003	0.003	0.004	0.003	0.003	0.002		
	(0.20)	(0.24)	(0.07)	(0.13)	(0.16)	(0.17)		
ΔLNIGPA		0.92		0.84		0.54		
		(0.01)		(0.04)		(0.22)		
ΔLNCLP		1.05		2.53		2.60		
		(0.20)		(0.02)		(0.00)		
AR(1)	0.36	0.32	0.31	0.29	0.27	0.26		
	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)		
Variance Equation								
Intercept	0.002	0.003	0.004	0.003	0.004	0.002		
	(0.31)	(0.17)	(0.46)	(0.36)	(0.26)	(0.26)		
ARCH	0.05	0.10	0.09	0.13	0.11	0.14		
	(0.22)	(0.13)	(0.03)	(0.18)	(0.07)	(0.05)		
GARCH	0.89	0.83	0.81	0.81	0.80	0.81		
	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)		
Model Information								
Obs	142	142	142	142	142	142		
Adj R ²	0.35	0.36	0.21	0.24	0.13	0.18		
AIC	-0.32	-0.32	-0.37	-0.43	- 0.49	-0.55		
Diagnostic Tests								
ARCH LM	2.25	1.79	0.70	0.59	0.81	0.66		
(12 lags)	(0.01)	(0.06)	(0.75)	(0.84)	(0.64)	(0.79)		
DW Stat	1.90	1.83	1.86	1.73	1.87	1.73		
JQB	14.27	5.76	5.38	2.87	4.02	1.85		
	(0.00)	(0.06)	(0.07)	(0.24)	(0.13)	(0.40)		

Table C2. GARCH (1,1) Model (with △PDBC30D and △COREYOY)

Note: All vars are in diff, *p*-values are in parenthesis